



# The price of unfairness in linear bandits with biased feedback

Solenne Gaucher<sup>(1)</sup>, Alexandra Carpentier<sup>(2)</sup> et Christophe Giraud<sup>(1)</sup>

- (1) Université Paris Saclay
- (2) Postdam Universität

London, June 2022

Fairness in Machine Learning: a major societal concern

Machine Learning is ubiquitous in daily life



Fairness in Machine Learning: a major societal concern

Machine Learning is ubiquitous in daily life

05-17-19

# Schools are using software to help pick who gets in. What could go wrong?

Admissions officers are increasingly turning to automation and AI with the hope of streamlining the application process and leveling the playing field.

Fairness in Machine Learning: a major societal concern

Machine Learning is ubiquitous in daily life

#### SCIENCE ADVANCES | RESEARCH ARTICLE

#### **RESEARCH METHODS**

# The accuracy, fairness, and limits of predicting recidivism

#### Julia Dressel and Hany Farid\*

Algorithms for predicting recidivism are commonly used to assess a criminal defendant's likelihood of committing a crime. These predictions are used in pretrial, parole, and sentencing decisions. Proponents of these systems argue that big data and advanced machine learning make these analyses more accurate and less biased than humans. We show, however, that the widely used commercial risk assessment software COMPAS is no more accurate or fair than predictions made by people with little or no criminal justice expertise. In addition, despite COMPAS's collection of 137 features, the same accuracy can be achieved with a simple linear predictor with only two features. Copyright © 2018 The Authors, some rights reserved; exclusive licensee American Association for the Advancement of Science. No claim to original U.S. Government Works. Distributed under a Creative Commons Attribution NonCommercial License 4.0 (CC BY-NC). A simple fairness problem in sequential decision making

## Sequential decision making with covariates

#### Actions and rewards

**Actions:** indexed by  $\mathcal{X} \subset \mathbb{R}^d$ **Reward:** f(x) for action  $x \in \mathcal{X}$  for some <u>unobserved</u>  $f : \mathcal{X} \to \mathbb{R}$ .

#### Bandit problems with covariates

At each round  $t = 1, \ldots, T$ 

- the agent chooses an action  $x_t \in \mathcal{X}$ , based on her historical data
- she <u>observes</u> the feedback  $y_t = f(x_t) + \xi_t$ , with  $(\xi_t)_{t \ge 1}$  independent.

#### Goal

Maximise the <u>unobserved</u> cumulated reward  $\sum_t f(x_t)$ .

## Sequential decision making with covariates

Optimal oracle strategy

Sample at each round  $x^* \in \underset{x \in \mathcal{X}}{\operatorname{argmax}} f(x)$ .

Infeasible since  $x^*$  is unknown...

### Decision maker objective

Minimize the regret

$$R_T = \mathbb{E}\Big[\sum_{t \leq T} (f(x^*) - f(x_t))\Big]$$

**Linear bandit:**  $f(x) = x^{\top} \theta^*$  with  $\theta^*$  unknown (very popular in applications)

## Fairness issue

#### Unfairness in ML

The main cause of unfairness in applications of ML algorithms, is the presence of biases in the data.

#### Setting

- Each action x is characterized by an (observed) attribute  $z_x \in \{-1, +1\}$  (e.g. gender)
- The feedbacks are biased depending on  $z_x$

#### Questions:

- What is the impact of such a bias?
- e How do handle it?

## **Biased Linear Bandits**

#### Linear bandit with biased feedbacks

At each round  $t = 1, \ldots, T$ 

- the agent chooses an action  $x_t\in\mathcal{X}\subset\mathbb{R}^d,$  whose sensitive attribute is  $z_{x_t}\in\{-1,1\}$
- she receives the unobserved reward  $x_t^{\top} \gamma^*$ ;
- she observes the biased feedback  $y_t = x_t^{\top} \gamma^* + \mathbf{z}_{\mathbf{x}_t} \omega^* + \xi_t$ .

#### Objective

Minimize the regret

$$R_T = \mathbb{E}\Big[\sum_{t \leq T} (x^* - x_t)^\top \gamma^*\Big], \quad ext{where} \quad x^* \in \operatorname*{argmax}_{x \in \mathcal{X}} x^\top \gamma^*.$$

## An insightful toy example

## Toy example - unbiased feedbacks



**Unbiased feedback:**  $\gamma^T x + \xi$  with  $\xi$  sub-Gaussian

Best action:

- if  $\gamma = \gamma_k$  action  $x_k$  is optimal, k = 1, 2;
- in both cases  $x_3$  is very suboptimal.

## Toy example - unbiased feedbacks



**①** The feedbacks differ by  $2\delta$  between  $x_1$  and  $x_2$ 

- Confidence intervals have width  $\propto \sqrt{\log(\text{confidence}^{-1})/N_{x_k}(t)}$ So:
  - if  $N_{x_k}(T) \leq \delta^{-2}$ , for k = 1 or 2: we cannot find the best action, and  $R_T = \Theta(T\delta)$ ;
  - if  $N_{x_k}(T) \gtrsim \delta^{-2} \log(T)$ , for k = 1 and 2: we find the best action with confidence 1/T, and  $R_T = \Theta(\delta \cdot \delta^{-2} \log(T)) = \Theta(\delta^{-1} \log(T))$ ;

Christophe Giraud (Orsay)

10 / 34

## Toy example - unbiased feedbacks

- if  $N_{x_k}(T) \le \delta^{-2}$ , for k = 1 or 2: we cannot find the best action, and  $R_T = \Theta(T\delta)$ ;
- if  $N_{x_k}(T) \ge \delta^{-2} \log(T)$ , for k = 1 and 2: we find the best action with confidence 1/T, and the regret is  $R_T = \Theta(\delta^{-1} \log(T))$ ;

#### Optimal regret with unbiased feedbacks

Large T regret: 
$$R_T = \Theta(\delta^{-1} \log(T))$$
, when  $T \to \infty$ ;

**Worst case regret:** the worst case is when  $\delta^{-2} = T$ , and then

$$R_T = \Theta(T\delta) = \Theta(\sqrt{T}).$$

## Toy example - biased feedbacks



**Biased feedback:**  $\gamma^T x + z_x \omega + \xi$  with  $\xi$  sub-Gaussian.

For  $(\gamma_1, \omega_1)$  and  $(\gamma_2, \omega_2)$ , the feedbacks are identical for  $x_1, x_2$ :

$$\begin{aligned} x_1^{\top} \gamma_1 + z_{x_1} \omega_1 &= x_1^{\top} \gamma_2 + z_{x_1} \omega_2 = 1, \\ \text{and} \quad x_2^{\top} \gamma_1 + z_{x_2} \omega_1 &= x_2^{\top} \gamma_2 + z_{x_2} \omega_2 = 1. \end{aligned}$$

 $\implies$  We need to sample the very sub-optimal action  $x_3$  to discriminate between  $(\gamma_1, \omega_1)$  and  $(\gamma_2, \omega_2)$ .

Christophe Giraud (Orsay)

## Toy example - biased feedbacks



The feedback when choosing  $x_3$  differs by  $4\delta$  between  $(\gamma_1, \omega_1)$  and  $(\gamma_2, \omega_2)$ :

- if  $N_{x_3}(T) \leq \delta^{-2}$ : we cannot find the best action, and  $R_T = \Theta(T\delta)$ ;
- if  $N_{x_3}(T) \gtrsim \delta^{-2} \log(T)$ : we find the best and  $R_T = \Theta(\delta^{-2} \log(T))$ .

Worst-case regret is achieved for  $\delta = T^{-1/3}$ , and  $R_T = \tilde{\Theta}(T^{2/3})$ .

## Toy example: summary

	Unbiased	Biased
Asymptotic regret	$R_T = \Theta(\delta^{-1}\log(T))$	$R_T = \Theta(\delta^{-2}\log(T))$
Worst case regret	for $\delta = T^{-1/2}$	for $\delta = T^{-1/3}$
	$R_T^* = \Theta(\sqrt{T})$	$R_T^* =  ilde{\Theta}(T^{2/3})$

Question: What is the price of biased feedbacks in general?

## Refresher on unbiased linear bandits

## A general recipe in bandits: successive elimination

#### Confidence interval

 $\mathcal{I}_x(t) :=$  confidence interval (at a prescribed level) for f(x) from the data collected up to time t.

#### Recipe

If "
$$\mathcal{I}_x(t) < \max_{x'} \mathcal{I}_{x'}(t)$$
", drop out the action  $x$  from  $\mathcal{X}$ .

If " $\mathcal{I}_x(t) \cap \max_{x'} \mathcal{I}_{x'}(t) 
eq \emptyset$ " then get more samples to shrink  $\mathcal{I}_x$ 

## Successive Elimination (principle)

 $\underline{\text{REPEAT}}$  for  $\varepsilon \searrow 0$ 

 $\bullet$  Sample a minimal number of actions from  ${\mathcal X}$  to get

$$|\mathcal{I}_x(t)| \leq arepsilon$$
 for all  $x \in \mathcal{X}$ ;

• Drop out from  $\mathcal X$  all actions x such that  $\mathcal I_x(t) < \max_{x'} \mathcal I_{x'}(t)$ .

## Linear bandits

#### Unbiased linear bandit

**Reward:**  $f(x) = x^{\top} \theta^*$ **Feedback:**  $y = x^{\top} \theta^* + \xi$  with  $\xi$  subGaussian(1).

#### Assumptions

$$|\mathcal{X}| = k < \infty$$
 and  $|x^{\top} \theta^*| \le 1$  for all  $x \in \mathcal{X}$ .

## Confidence bounds

#### OLS estimator

For *n* sampled actions  $x_1, ..., x_n$  in  $\mathcal{X}$ , the OLS estimator is

$$\widehat{\theta} = V^+ \sum_{s \leq n} x_s y_s, \quad \text{where} \quad V = \sum_{s \leq n} x_s x_s^\top,$$

and  $V^+$  is the Moore-Penrose pseudo inverse.

#### Confidence bound

If  $x_1, ..., x_n$  are fixed, then for all  $x \in \text{Range}(V)$ ,

$$\mathbb{P}\left(\left|\left(\widehat{\theta}-\theta^*\right)^\top x\right| \le \sqrt{2 \left\|x\right\|_{V^+}^2 \log\left(\frac{1}{\delta}\right)}\right) \ge 1-\delta.$$

where  $||x||_{V^+}^2 := x^\top V^+ x$ .

## Confidence bounds

#### OLS estimator

For *n* sampled actions  $x_1, ..., x_n$  in  $\mathcal{X}$ , the OLS estimator is

$$\widehat{\theta} = V^+ \sum_{s \leq n} x_s y_s, \quad \text{where} \quad V = \sum_{s \leq n} x_s x_s^\top,$$

and  $V^+$  is the Moore-Penrose pseudo inverse.

#### Confidence bound

If  $x_1, ..., x_n$  are fixed, then for all  $x \in \text{Range}(V)$ ,

$$\mathbb{P}\left(\left|\left(\widehat{\theta}-\theta^*\right)^\top x\right| \leq \sqrt{2 \left\|x\right\|_{V^+}^2 \log\left(\frac{1}{\delta}\right)}\right) \geq 1-\delta.$$

where  $||x||_{V^+}^2 := x^\top V^+ x$ .

## G-optimal design

If we choose each action x exactly  $\mu(x)$  times

$$\sum_{s\leq n} x_s x_s^\top = V(\mu) := \sum_{x\in\mathcal{X}} \mu(x) x x^\top$$

G-optimal design

$$\mu_n^* \in \underset{|\mu|=n}{\operatorname{argmin}} \max_{x \in \mathcal{X}} \|x\|_{V(\mu)^+}^2 . \quad \text{(G-optimal design)}$$

fulfills

$$\max_{\mathbf{x}\in\mathcal{X}} \|\mathbf{x}\|_{V(\mu_n^*)^+}^2 \leq \frac{d}{n}.$$

#### Confidence bound

If each action  $x \in \mathcal{X}$  is sampled  $\mu_n^*(x)$  times with  $n = \frac{2d}{\epsilon^2} \log \left(k\delta^{-1}\right)$ , then  $\max_{x \in \mathcal{X}} \left| \left(\widehat{\theta} - \theta^*\right)^\top x \right| \le \epsilon$ , with probability at least  $1 - \delta$ .

Christophe Giraud (Orsay)

## G-optimal design

If we choose each action x exactly  $\mu(x)$  times

$$\sum_{s\leq n} x_s x_s^{\top} = V(\mu) := \sum_{x\in\mathcal{X}} \mu(x) x x^{\top}$$

G-optimal design

$$\mu_n^* \in \operatorname*{argmin}_{|\mu|=n} \; \max_{x \in \mathcal{X}} \|x\|_{V(\mu)^+}^2 \; . \quad \text{(G-optimal design)}$$

fulfills

$$\max_{x\in\mathcal{X}} \|x\|_{V(\mu_n^*)^+}^2 \leq \frac{d}{n}.$$

#### Confidence bound

If each action  $x \in \mathcal{X}$  is sampled  $\mu_n^*(x)$  times with  $n = \frac{2d}{\epsilon^2} \log(k\delta^{-1})$ , then  $\max_{x \in \mathcal{X}} \left| \left( \widehat{\theta} - \theta^* \right)^\top x \right| \le \epsilon$ , with probability at least  $1 - \delta$ .

Christophe Giraud (Orsay)

## G-optimal design

If we choose each action x exactly  $\mu(x)$  times

$$\sum_{s\leq n} x_s x_s^{\top} = V(\mu) := \sum_{x\in\mathcal{X}} \mu(x) x x^{\top}$$

G-optimal design

$$\mu_n^* \in \operatorname*{argmin}_{|\mu|=n} \; \max_{x \in \mathcal{X}} \|x\|_{V(\mu)^+}^2 \; . \quad \text{(G-optimal design)}$$

fulfills

$$\max_{x\in\mathcal{X}} \|x\|_{V(\mu_n^*)^+}^2 \leq \frac{d}{n}.$$

#### Confidence bound

If each action  $x \in \mathcal{X}$  is sampled  $\mu_n^*(x)$  times with  $n = \frac{2d}{\epsilon^2} \log \left(k\delta^{-1}\right)$ , then  $\max_{x \in \mathcal{X}} \left| \left(\widehat{\theta} - \theta^*\right)^\top x \right| \le \epsilon, \quad \text{with probability at least } 1 - \delta.$ 

## Phased Elimination algorithm

#### PHASED ELIMINATION (Lattimore and Szepesvári, 2020)

Input 
$$\mathcal{X}_{1} = \mathcal{X}$$
.  
For  $l = 1, 2, ...$   
•  $\epsilon_{l} \leftarrow 2^{-l}, \quad n_{l} \leftarrow \frac{2d}{\epsilon_{l}^{2}} \log (kl(l+1)T)$   
•  $\mathcal{X}_{l+1} \leftarrow \text{G-EXPLORE-AND-ELIMINATE}(\mathcal{X}_{l}, n_{l}, \epsilon_{l})$   
End For

### G-EXPLORE-AND-ELIMINATE( $\mathcal{X}_{l}, n_{l}, \epsilon_{l}$ )

- sample  $\mu_{n_l}^*(x)$  times each action  $x \in \mathcal{X}_l$
- compute OLS estimator  $\widehat{\theta}$

**Return** 
$$\mathcal{X}_l \setminus \left\{ x \in \mathcal{X}_l : x^\top \widehat{\theta} + \epsilon_l \leq \max_{x \in \mathcal{X}_l} x^\top \widehat{\theta} - \epsilon_l \right\}$$

## Phased Elimination algorithm

#### PHASED ELIMINATION (Lattimore and Szepesvári, 2020)

Input 
$$\mathcal{X}_{1} = \mathcal{X}$$
.  
For  $l = 1, 2, ...$   
•  $\epsilon_{l} \leftarrow 2^{-l}, n_{l} \leftarrow \frac{2d}{\epsilon_{l}^{2}} \log (kl(l+1)T)$   
•  $\mathcal{X}_{l+1} \leftarrow \text{G-EXPLORE-AND-ELIMINATE}(\mathcal{X}_{l}, n_{l}, \epsilon_{l})$   
End For

#### G-EXPLORE-AND-ELIMINATE( $\mathcal{X}_l, n_l, \epsilon_l$ )

- sample  $\mu_{n_l}^*(x)$  times each action  $x \in \mathcal{X}_l$
- compute OLS estimator  $\widehat{\theta}$

**Return** 
$$\mathcal{X}_{I} \setminus \left\{ x \in \mathcal{X}_{I} : x^{\top} \widehat{\theta} + \epsilon_{I} \leq \max_{x \in \mathcal{X}_{I}} x^{\top} \widehat{\theta} - \epsilon_{I} \right\}$$

## Phased Elimination algorithm

#### Gaps

**Gaps:**  $\Delta_x = (x^* - x)^{\top} \theta^*$ **Minimal gap:**  $\Delta_{\min} = \min \{\Delta_x : \Delta_x > 0\}$ 

#### Theorem

**Asymptotic regret:**  $R_T \leq d\Delta_{\min}^{-1} \log(T)$ 

The worst case regret:  $R_T^* \lesssim C \sqrt{dT \log(kT)}$ .

## **Biased linear bandits**

## **Biased Linear Bandits**

#### Our setting

At each round  $t = 1, \ldots, T$ 

- the agent chooses an action  $x_t\in\mathcal{X}\subset\mathbb{R}^d$  , whose sensitive attribute is  $z_{x_t}\in\{-1,1\}$
- she receives the unobserved reward  $x_t^{\top} \gamma^*$ ;
- she observes the biased feedback

$$\begin{split} \chi_t &= x_t^{\top} \gamma^* + \mathbf{z}_{\mathbf{x}_t} \omega^* + \xi_t \\ &= \mathbf{a}_{\mathbf{x}_t}^{\top} \theta^* + \xi_t, \qquad \text{where } \mathbf{a}_{\mathbf{x}} = \begin{pmatrix} \mathbf{x} \\ \mathbf{z}_{\mathbf{x}} \end{pmatrix}, \ \theta^* = \begin{pmatrix} \gamma^* \\ \omega^* \end{pmatrix} \end{split}$$

#### Objective

Minimize the regret 
$$R_T = \mathbb{E}\Big[\sum_{t \leq T} (x^* - x_t)^\top \gamma^*\Big]$$
, where  $x^* \in \underset{x \in \mathcal{X}}{\operatorname{argmax}} x^\top \gamma^*$ .

## **Biased Linear Bandits**

### Our setting

At each round  $t = 1, \ldots, T$ 

- the agent chooses an action  $x_t \in \mathcal{X} \subset \mathbb{R}^d$ , whose sensitive attribute is  $z_{x_t} \in \{-1,1\}$
- she receives the unobserved reward  $x_t^{\top} \gamma^*$ ;
- she observes the biased feedback

$$\begin{split} \varphi_t &= x_t^{\top} \gamma^* + \mathbf{z}_{\mathbf{x}_t} \omega^* + \xi_t \\ &= \mathbf{a}_{\mathbf{x}_t}^{\top} \theta^* + \xi_t, \qquad \text{where } \mathbf{a}_{\mathbf{x}} = \begin{pmatrix} \mathbf{x} \\ \mathbf{z}_{\mathbf{x}} \end{pmatrix}, \ \theta^* = \begin{pmatrix} \gamma^* \\ \omega^* \end{pmatrix} \end{split}$$

#### Objective

Minimize the regret 
$$R_T = \mathbb{E}\Big[\sum_{t \leq T} (x^* - x_t)^\top \gamma^*\Big]$$
, where  $x^* \in \underset{x \in \mathcal{X}}{\operatorname{argmax}} x^\top \gamma^*$ .

Christophe Giraud (Orsay)

## A first naive idea

#### Bias issue

We have the linear model  $y = a_x^{\top} \theta^* + \xi$ , but the reward is

$$x^{\top}\gamma^* = a_x^{\top}\theta^* - \mathbf{z}_x\omega^*.$$

#### Naive DEBIASED G-EXPLORE-AND-ELIMINATE( $\mathcal{X}_l, n_l, \epsilon_l$ )

- $\mu_{n_l}^* \leftarrow \text{G-optimal design}(\{a_x : x \in \mathcal{X}_l\}, n_l)$
- sample  $\mu_{n_l}^*(x)$  times each action  $x \in \mathcal{X}_l$
- compute OLS estimator  $\widehat{\theta} = \begin{pmatrix} \widehat{\gamma} \\ \widehat{\omega} \end{pmatrix}$

$$\textbf{Return } \mathcal{X}_l \setminus \left\{ x \in \mathcal{X}_l : x^\top \widehat{\gamma} + \epsilon_l \leq \max_{x \in \mathcal{X}_l} x^\top \widehat{\gamma} - \epsilon_l \right\}$$

## A first naive idea

#### Bias issue

We have the linear model  $y = a_x^{\top} \theta^* + \xi$ , but the reward is

$$x^{\top}\gamma^* = a_x^{\top}\theta^* - \mathbf{z}_x\omega^*.$$

#### Naive DEBIASED G-EXPLORE-AND-ELIMINATE( $\mathcal{X}_{l}, n_{l}, \epsilon_{l}$ )

- $\mu_{n_l}^* \leftarrow \text{G-optimal design}(\{a_x : x \in \mathcal{X}_l\}, n_l)$
- sample  $\mu_{n_l}^*(x)$  times each action  $x \in \mathcal{X}_l$
- compute OLS estimator  $\hat{\theta} = \begin{pmatrix} \hat{\gamma} \\ \hat{\omega} \end{pmatrix}$

$$\textbf{Return } \mathcal{X}_l \setminus \left\{ x \in \mathcal{X}_l : x^\top \widehat{\gamma} + \epsilon_l \leq \max_{x \in \mathcal{X}_l} x^\top \widehat{\gamma} - \epsilon_l \right\}$$

## Failure of the naive idea

#### Caveat

We have with probability at least  $1-\delta$ 

$$\max_{\boldsymbol{x}\in\mathcal{X}_I}\left|\left(\widehat{\theta}-\theta^*\right)^\top \boldsymbol{a}_{\boldsymbol{x}}\right|\leq\epsilon_I.$$

But, we have no control on the reward  $\bigcirc$ 

$$\max_{x \in \mathcal{X}_I} \left| (\widehat{\gamma} - \gamma^*)^\top x \right| \le ??$$

#### Remedy

We need an additional step of optimal-designed estimation of the bias  $\omega^*.$ 

## Failure of the naive idea

#### Caveat

We have with probability at least  $1-\delta$ 

$$\max_{\boldsymbol{x}\in\mathcal{X}_I}\left|\left(\widehat{\theta}-\theta^*\right)^\top \boldsymbol{a}_{\boldsymbol{x}}\right|\leq\epsilon_I.$$

But, we have no control on the reward  $\bigcirc$ 

$$\max_{x \in \mathcal{X}_l} \left| (\widehat{\gamma} - \gamma^*)^\top x \right| \le ??$$

#### Remedy

We need an additional step of optimal-designed estimation of the bias  $\omega^{\ast}.$ 

## **Bias estimation**

OLS estimation (best unbiased linear estimator)

• Sample  $\mu(x)$  times each  $x \in \mathcal{X}$ 

• Compute the OLS estimator 
$$\widehat{ heta} = egin{pmatrix} \widehat{\gamma} \\ \widehat{\omega} \end{pmatrix}$$

Then,

$$\mathbb{P}\left(\left|\widehat{\omega}-\omega^*\right| \leq \sqrt{2 \left\| \boldsymbol{e_{d+1}} \right\|_{\boldsymbol{V}(\mu)^+}^2 \log\left(1/\delta\right)}\right) \geq 1-\delta_{\boldsymbol{e_{d+1}}}$$

where 
$$V(\mu) = \sum_x \mu(x) a_x a_x^ op$$
 and  $e_{d+1} = (0, \dots, 0, 1).$ 

#### Regret for bias estimation

When we sample  $\mu(x)$  times each  $x \in \mathcal{X}$ , we suffer the regret

$$\sum_{\mathbf{x}\in\mathcal{X}}\mu(\mathbf{x})\Delta_{\mathbf{x}}, \quad ext{where} \quad \Delta_{\mathbf{x}}=(\mathbf{x}^*-\mathbf{x})^{ op}\gamma^*.$$

## **Bias estimation**

OLS estimation (best unbiased linear estimator)

• Sample  $\mu(x)$  times each  $x \in \mathcal{X}$ 

• Compute the OLS estimator 
$$\widehat{ heta} = egin{pmatrix} \widehat{\gamma} \\ \widehat{\omega} \end{pmatrix}$$

Then,

$$\mathbb{P}\left(\left|\widehat{\omega}-\omega^*\right| \leq \sqrt{2 \left\|\boldsymbol{e_{d+1}}\right\|_{\boldsymbol{V}(\mu)^+}^2 \log\left(1/\delta\right)}\right) \geq 1-\delta_{\boldsymbol{e_{d+1}}}$$

where 
$$V(\mu) = \sum_x \mu(x) a_x a_x^\top$$
 and  $e_{d+1} = (0, \dots, 0, 1)$ .

#### Regret for bias estimation

When we sample  $\mu(x)$  times each  $x \in \mathcal{X}$ , we suffer the regret

$$\sum_{\mathbf{x}\in\mathcal{X}}\mu(\mathbf{x})\Delta_{\mathbf{x}}, \quad ext{where} \ \ \Delta_{\mathbf{x}}=(\mathbf{x}^*-\mathbf{x})^{ op}\gamma^*.$$

## $\Delta$ -optimal design

#### $\Delta\text{-optimal}$ design

For  $\Delta = (\Delta_x)_{x \in \mathcal{X}}$ , we introduce

$$\mu^{\Delta} = \underset{\substack{\mu \in \mathcal{M}_{e_{d+1}}^{\mathcal{X}} \text{ s.t.} \\ \|e_{d+1}\|_{V(\mu)^{+}}^{\mathcal{X}} \leq 1}}{\operatorname{argmin}} \sum_{x \in \mathcal{X}} \mu(x) \Delta_{x} \qquad (\Delta \text{-optimal design})$$

## $\Delta\text{-}optimal$ design

#### $\Delta\text{-optimal}$ design

For  $\Delta = (\Delta_x)_{x \in \mathcal{X}}$ , we introduce

$$\sigma^{-1}\mu^{\Delta} = \operatorname{argmin}_{\substack{\mu \in \mathcal{M}_{e_{d+1}}^{\mathcal{X}} \text{ s.t.} \\ \|e_{d+1}\|_{V(\mu)^{+}}^{\mathcal{X}} \leq \sigma^{2}}} \sum_{x \in \mathcal{X}} \mu(x)\Delta_{x} \quad (\Delta \text{-optimal design})$$

## Regret for $\Delta$ -optimal bias estimation

#### Regret for bias estimation

The minimal regret for estimating the bias  $\omega^*$  with precision  $\epsilon$  and confidence  $1-\delta$  is

minimal regret for bias estimation 
$$=rac{2\kappa(\Delta)\log(\delta^{-1})}{\epsilon^2}$$
 .

where

$$\kappa(\Delta) = \sum_{x \in \mathcal{X}} \mu^{\Delta}(x) \Delta_x$$

characterizes the difficulty of bias estimation in our setting.

In practice

In practice,  $\Delta_{\chi}$  is unknown, we have to rely on some (upper) estimates  $\overline{\Delta}_{\chi}.$ 

## Regret for $\Delta$ -optimal bias estimation

#### Regret for bias estimation

The minimal regret for estimating the bias  $\omega^*$  with precision  $\epsilon$  and confidence  $1-\delta$  is

minimal regret for bias estimation 
$$=rac{2\kappa(\Delta)\log(\delta^{-1})}{\epsilon^2}$$
 .

where

$$\kappa(\Delta) = \sum_{x \in \mathcal{X}} \mu^{\Delta}(x) \Delta_x$$

characterizes the difficulty of bias estimation in our setting.

In practice

In practice,  $\Delta_x$  is unknown, we have to rely on some (upper) estimates  $\overline{\Delta}_x$ .

Christophe Giraud (Orsay)

## FAIR PHASED ELIMINATION

#### 3 main ingredients

- Actions can be compared within a group: we apply G-EXPLORE-AND-ELIMINATE within a group
- 2 bias correction based on  $\widehat{\Delta}$ -optimal design

3 bias estimation breaking criterion: to avoid a too high regret

## FAIR PHASED ELIMINATION algorithm

#### FAIR PHASED ELIMINATION

$nput\ \mathcal{Z} = \{-1, +1\}, \ and\ \mathcal{X}_1^z = \{x \in \mathcal{X} : z_x = z\} \ for\ z \in \mathcal{Z}$
For $l = 1, 2,$
• $\epsilon_{l} \leftarrow 2^{-l}, \ n_{l} \leftarrow \frac{2(d+1)}{\epsilon_{l}^{2}} \log\left(\frac{kl(l+1)}{\delta}\right), \ m_{l} \leftarrow \frac{2}{\epsilon_{l}^{2}} \log\left(\frac{l(l+1)}{\delta}\right)$
• For $z \in \mathcal{Z}$
$   \mathcal{X}_{l+1}^{z}, \widehat{\theta}_{l}^{(z)} \leftarrow \text{G-Explore-And-Eliminate}(\mathcal{X}_{l}^{z}, \textit{n}_{l}, \epsilon_{l}) $
• If $Z = \{-1, +1\}$
• If $\epsilon_l \leq \left(\kappa(\widehat{\Delta}^l)\log(\mathcal{T})/\mathcal{T}\right)^{1/3}$ , then break and sample best empirical
action for remaining time
$\widehat{\omega}_{I} \leftarrow \Delta \text{-} \text{Explore}(\widehat{\Delta}', m_{I})$
$\blacktriangleright \ \widehat{m}_x \leftarrow a_x^\top \widehat{\theta}_I^{(z)} - z \widehat{\omega}_I \text{ for } x \in \mathcal{X}_I^{(-1)} \cup \mathcal{X}_I^{(1)}, \text{ update } \widehat{\Delta}^I$
► If $\exists z \in \mathcal{Z}$ s.t. $\max_{x \in \mathcal{X}_l^{(z)}} \widehat{m}_x \ge \max_{x \in \mathcal{X}_l^{(-z)}} \widehat{m}_x + 4\epsilon_l$ then $\mathcal{Z} \leftarrow \{z\}$

## **Optimal Regret for Biased Linear Bandits**

## Geometry of worst-case regret

Theorem (Gaucher, Carpentier, Giraud, 2022)

Define

$$\kappa_* = \min_{\substack{\mu \in \mathcal{M}_{e_{d+1}}^{\mathcal{X}} \ s.t. \\ \|e_{d+1}\|_{V(\mu)^+}^2 \le 1}} \max_{\substack{ heta: |\mathbf{a}_x^ op \theta| \le 1 \ x \in \mathcal{X}}} \mu(x) \Delta_x( heta).$$

Then, FAIR PHASED ELIMINATION algorithm fulfills

$$R_T \leq C \, \kappa_*^{1/3} \, T^{2/3} \log(T)^{1/3}, \qquad for \quad T \geq T_{k,d,\kappa_*}.$$

#### Remarks

• Matching lower bound up to a  $\log(T)^{1/3}$ ;

•  $\kappa_*^{1/3}$  captures the dependency on the geometry of the set of actions;

• Regret in  $\tilde{\Theta}(T^{2/3})$  instead of  $\tilde{\Theta}(T^{1/2})$  is the price for debiasing the rewards.

## Geometry of worst-case regret

Theorem (Gaucher, Carpentier, Giraud, 2022)

Define

$$\kappa_* = \min_{\substack{\mu \in \mathcal{M}_{e_{d+1}}^{\mathcal{X}} \text{ s.t.} \\ \|e_{d+1}\|_{V(\mu)^+}^2 \leq 1}} \max_{ heta: |a_x^ op \theta| \leq 1} \sum_{x \in \mathcal{X}} \mu(x) \Delta_x( heta).$$

Then, FAIR PHASED ELIMINATION algorithm fulfills

$$R_T \leq C \, \kappa_*^{1/3} \, T^{2/3} \log(T)^{1/3}, \qquad for \quad T \geq T_{k,d,\kappa_*}.$$

#### Remarks

• Matching lower bound up to a  $\log(T)^{1/3}$ ;

•  $\kappa_*^{1/3}$  captures the dependency on the geometry of the set of actions;

• Regret in  $\tilde{\Theta}(T^{2/3})$  instead of  $\tilde{\Theta}(T^{1/2})$  is the price for debiasing the rewards.

## Geometry of worst-case regret

Theorem (Gaucher, Carpentier, Giraud, 2022)

Define

$$\kappa_* = \min_{\substack{\mu \in \mathcal{M}_{e_{d+1}}^{\mathcal{X}} \text{ s.t.} \\ \|e_{d+1}\|_{V(\mu)^+}^2 \leq 1}} \max_{ heta: |a_x^ op heta| \leq 1} \sum_{x \in \mathcal{X}} \mu(x) \Delta_x( heta).$$

Then, FAIR PHASED ELIMINATION algorithm fulfills

$$R_T \leq C \, \kappa_*^{1/3} \, T^{2/3} \log(T)^{1/3}, \qquad for \quad T \geq T_{k,d,\kappa_*}.$$

#### Remarks

- Matching lower bound up to a log(T)<sup>1/3</sup>;
- $\kappa_*^{1/3}$  captures the dependency on the geometry of the set of actions;
- Regret in  $\tilde{\Theta}(T^{2/3})$  instead of  $\tilde{\Theta}(T^{1/2})$  is the price for debiasing the rewards.

## Geometry of bias estimation



#### Lemma

$$\kappa^* = \Delta_{\max} \left(\frac{R+\delta}{R-\delta}\right)^2$$

with the largest  $\delta/R \in [0,1]$  such that a  $\delta$ -separation as above exists

Christophe Giraud (Orsay)

## $\Delta$ -dependent regret bound

$$\Delta_{\min} := \min_{x 
eq x^*} (x^* - x)^{ op} \gamma^* \qquad ext{and} \qquad \Delta_{
eq} := \min_{x 
eq x^*} (x^* - x)^{ op} \gamma^*.$$

Theorem (Gaucher, Carpentier, Giraud, 2022)

FAIR PHASED ELIMINATION algorithm fulfills

$$R_T \leq C\left(rac{d}{\Delta_{\min}} + rac{\kappa(\Delta ee \Delta_{
eq} ee arepsilon_{ au})}{\Delta_{
eq}^2}
ight)\log(T), \qquad for \quad T \geq k \lor e^{d\Delta_{\min}}$$
 ere  $arepsilon_T = (\kappa_*\log(T)/T)^{1/3}.$ 

#### Comments

• Some matching lower bounds;

d log(T)/Δ<sub>min</sub> is the (worst gap-dependent) regret of the classical linear bandit;
 κ(Δ) log(T)/Δ<sup>2</sup><sub>≠</sub> is the price for debiasing the rewards.

## $\Delta$ -dependent regret bound

$$\Delta_{\min} := \min_{x \neq x^*} (x^* - x)^\top \gamma^* \qquad \text{and} \qquad \Delta_{\neq} := \min_{z_x \neq z_{x^*}} (x^* - x)^\top \gamma^*.$$

Theorem (Gaucher, Carpentier, Giraud, 2022)

FAIR PHASED ELIMINATION algorithm fulfills

$$R_{\mathcal{T}} \leq C\left(\frac{d}{\Delta_{\min}} + \frac{\kappa \big(\Delta \lor \Delta_{\neq} \lor \varepsilon_{\mathcal{T}}\big)}{\Delta_{\neq}^2}\right) \log(\mathcal{T}), \qquad for \quad \mathcal{T} \geq k \lor e^{d\Delta_{\min}}$$

where  $\varepsilon_T = (\kappa_* \log(T)/T)^{1/3}$ .

#### Comments

• Some matching lower bounds;

 <sup>d log(T)</sup>/<sub>Δ<sub>min</sub></sub> is the (worst gap-dependent) regret of the classical linear bandit;

 <sup>κ(Δ) log(T)</sup>/<sub>Δ<sup>2</sup><sub>≠</sub></sub> is the price for debiasing the rewards.

## $\Delta$ -dependent regret bound

$$\Delta_{\min} := \min_{x \neq x^*} (x^* - x)^\top \gamma^* \qquad \text{and} \qquad \Delta_{\neq} := \min_{z_x \neq z_{x^*}} (x^* - x)^\top \gamma^*.$$

Theorem (Gaucher, Carpentier, Giraud, 2022)

FAIR PHASED ELIMINATION algorithm fulfills

$$R_{\mathcal{T}} \leq C\left(\frac{d}{\Delta_{\min}} + \frac{\kappa(\Delta \lor \Delta_{\neq} \lor \varepsilon_{\mathcal{T}})}{\Delta_{\neq}^2}\right) \log(\mathcal{T}), \qquad for \quad \mathcal{T} \geq k \lor e^{d\Delta_{\min}}$$

where  $\varepsilon_T = (\kappa_* \log(T)/T)^{1/3}$ .

#### Comments

- Some matching lower bounds;
- $\frac{d \log(T)}{\Delta_{\min}}$  is the (worst gap-dependent) regret of the classical linear bandit;
- $\frac{\kappa(\Delta)\log(T)}{\Delta^2_{\ell}}$  is the price for debiasing the rewards.

## Take home message

#### In biased linear bandit problems

• In the worst case, the regret can be  $\tilde{\Theta}(T^{2/3})$  instead of  $\tilde{\Theta}(\sqrt{T})$ . The geometric dependence is captured by the largest  $\delta$ -separation.

• In gap-depend worst case:  
• an additional 
$$\frac{\kappa(\Delta)\log(T)}{\Delta_{\neq}^2}$$
 term shows up  
• can be as easy as classical bandit if  $\frac{\kappa(\Delta)\log(T)}{\Delta_{\neq}^2} \leq \frac{d\log(T)}{\Delta_{\min}}$